

$$\Pr_{Y \sim \Sigma} \Pr(A_m(\zeta)) \xrightarrow{n \rightarrow \infty} 0 \quad (\text{convergence in probability})$$

$\left(\sum_m P(A_m(\zeta)) < \infty \right) \rightarrow a.s.$ $\sum_m A_m \leq \infty$
 \downarrow $\lim_{m \rightarrow \infty} \sum_m A_m = 0$
 $\Pr(A_m(\zeta)) \xrightarrow{n \rightarrow \infty} 0$ strongly

B-C Lemma $\Pr_{Y \sim \Sigma} \sum_m A_m \leq \infty \rightarrow a.s. X_m \rightarrow X$ as sample size (sufficient)

① $\sum_m P(A_m(\zeta)) < \infty$

② $\sum_m P(A_m(\zeta)) < \infty \rightarrow P(A_m(\zeta)) \xrightarrow{n \rightarrow \infty} 0$

$\Pr(X_m - X > \varepsilon) \rightarrow 0$ a.s.

$\left(\sum_m \frac{1}{m} < \infty \right) \rightarrow a.s. (X)$

$= \frac{1}{m^2} (0) \quad \sum_m \frac{1}{m^2} < \infty$

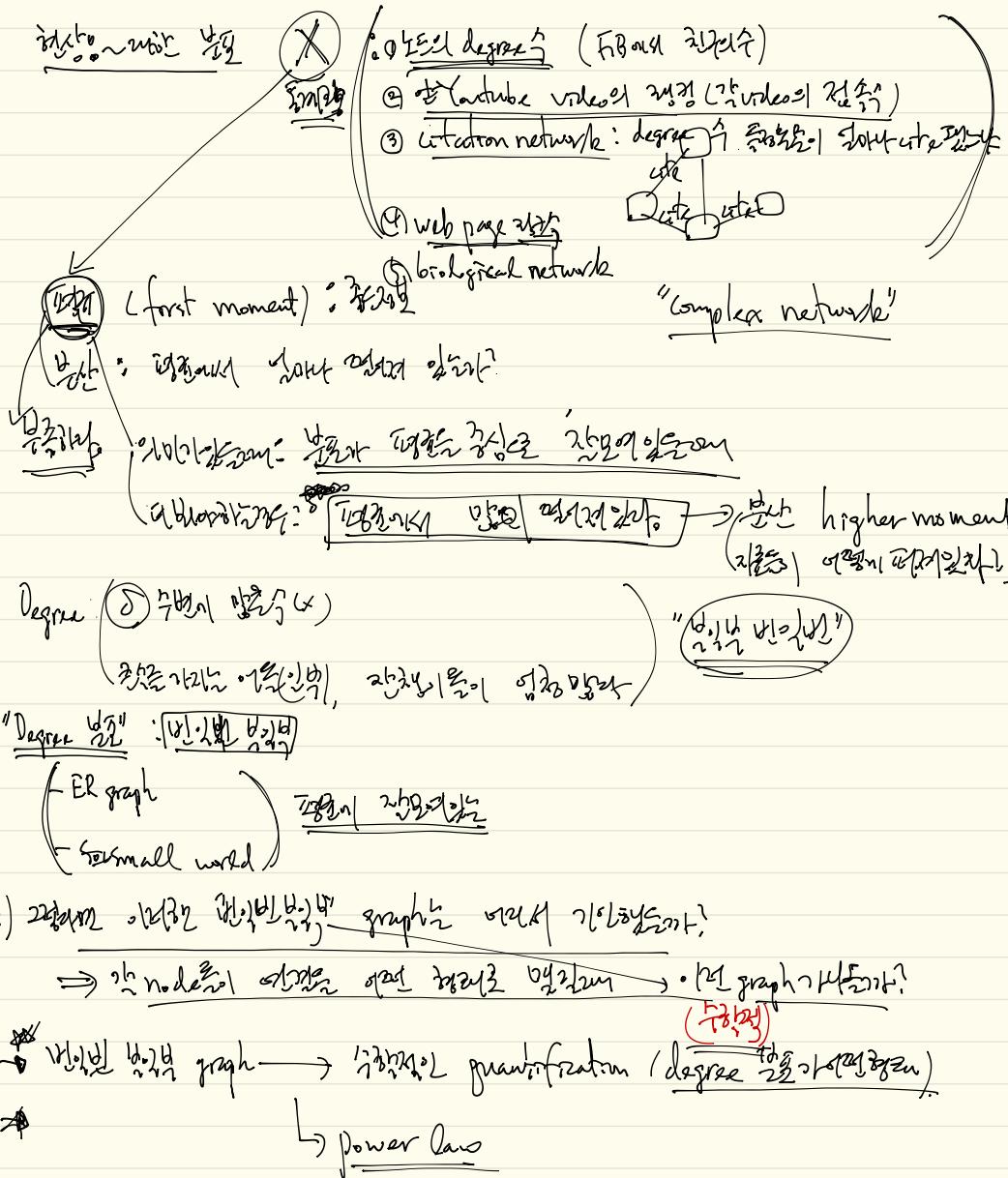
Gronwall's Lemma $\Pr_{Y \sim \Sigma} \sum_m A_m \leq \infty \rightarrow a.s. (X)$

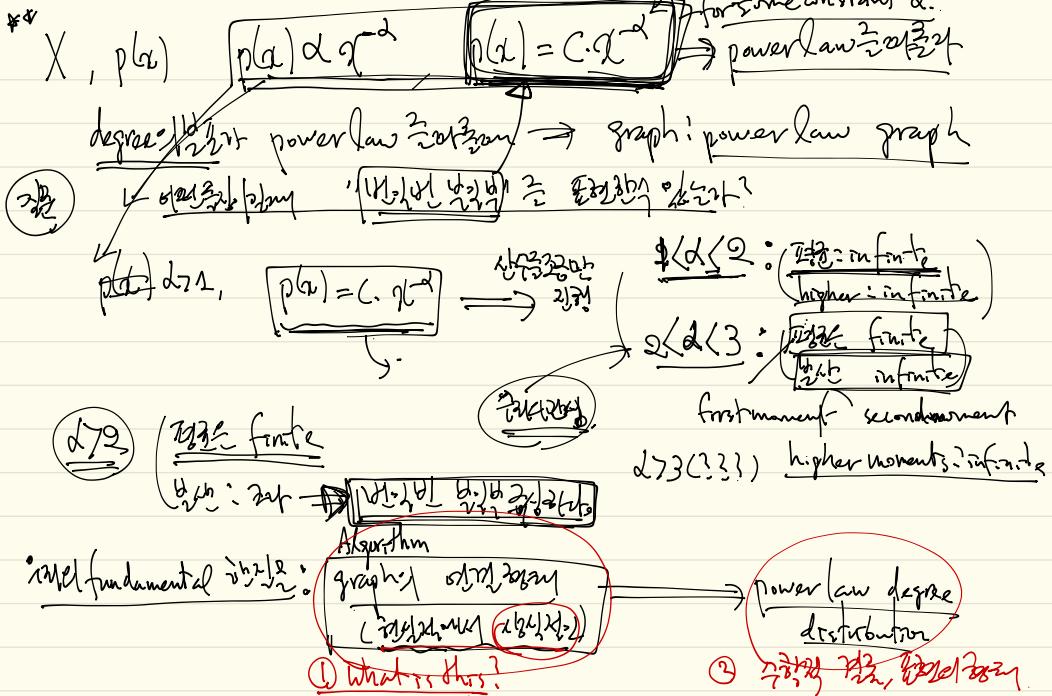
Mode of Convergence

$\Pr_{Y \sim \Sigma} \sum_m A_m \leq \infty \rightarrow a.s. (X)$

B.C test

Lecture 12 (chapter 7) Power laws via Preferential attachment





Barabasi \rightarrow algorithm 20102 [Preferential attachment] \Rightarrow power law

Preferential attachment (long 20102) (page 78 of 100)

- A graph is grown over time $G_0, G_1, G_2, \dots, G_t$ $\frac{G_t = (N(t), E(t))}{\text{random}}$
- every step t , we add one new node randomly
- Given G_t ($E(t) = (N(t), P(t))$), a new node v_{t+1} is added
 - ① with probability $\frac{1}{|E|}$ uniformly at random
 - ② with probability $\frac{d(v)}{|E|}$, a node $v \in N(t)$ is selected with $\frac{d(v)}{|E|}$ \rightarrow degree $d(v)$ \propto $\frac{1}{|E|}$

(Q) PA

power law degree distribution

(2nd) real research problem:

① What is the power law?

one solution

multiple solutions

weak
strong

probability distribution

$\pi_j \rightarrow \pi_i$ (rich-gets-rich)

America's At. inequality
Coupling

Yule process

better

1925

1949

1951

power law distribution

heavy tail,
distribution

scale-free network

(i) scale-free (Scale-invariant)

$\propto d^{-r}$

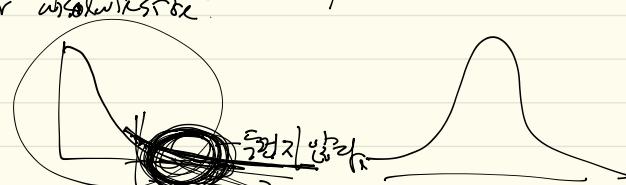
$P(d) \propto d^{-r}$: degree distribution \sim power law

$$\frac{P(d)}{P(d')} = \frac{C d^{-r}}{C d'^{-r}} = \left(\frac{d}{d'}\right)^{-r} \Rightarrow d \mapsto d^{\frac{1}{r}} \text{ rescale by any factor}$$

$d \leftrightarrow d'$:
 $d \downarrow$
and $d \downarrow : 10^3 / 10^2 = 10^{3-2} = 10^1$

"relative probabilities of different degrees depend only
on their ratios not their absolute size"

(ii) (heavy-tail) (tailored)
long-tail



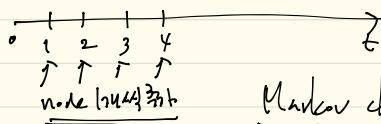
The distribution of RV X with distribution function F is said to have "heavy tail" if

$$\lim_{x \rightarrow \infty} e^{\lambda x} F(x) = \infty \quad \text{for all } \lambda > 0$$

$\Rightarrow \lim_{x \rightarrow \infty} e^{\lambda x} F(x) = \infty$

Power-law object \rightarrow heavy tail

$$e^{kx} P(X \geq x) = e^{kx} \int_x^{\infty} c \cdot t^{-\beta} dt = c e^{kx} \left(\frac{1}{\beta+1} \right)_x^{\infty} = c * e^{kx} \left(-\frac{x^{\beta+1}}{\beta+1} \right) : n \rightarrow \infty$$



Markov chain (discrete time)

$$\vec{X}(t) = [X_i(t)]$$

any $X_i(t)$: $X_i(t) \triangleq$ # of times degree i visited node i

$$P(X_i(t+1) = X_i(t) + 1 \mid \vec{X}(t)) = \frac{N(t)}{N(t)} + (1-\alpha) \frac{(i-1) \cdot X_i(t)}{2E(t)} \quad \text{with node } i$$

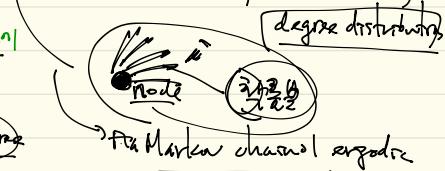
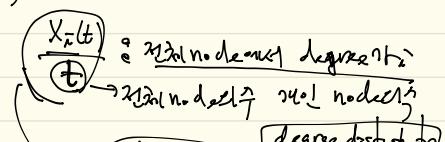
$$P(X_i(t+1) = X_i(t) \mid \vec{X}(t)) = \frac{N(t)}{N(t)} + (1-\alpha) \frac{i \cdot X_i(t)}{2E(t)} \quad \text{degree } i \text{ is } \text{random and } \text{deg} = \text{degree}$$

$$P(X_i(t+1) > X_i(t) \mid \vec{X}(t))$$

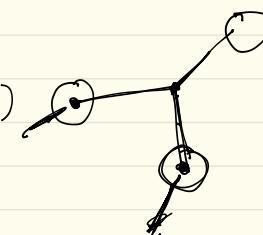
$$= 1 - ① - ②$$

$$i=1 \text{ or } 2 \text{ or } 3 \text{ or } 4$$

$$P(X_i(t+1) = X_i(t) \mid \vec{X}(t)) = 0 \quad (x)$$



for Markov chain ergodic
time average
ensemble average



$$P(X_i(t+1) = X_i(t) \mid \vec{X}(t))$$

$$= \alpha \frac{X_i(t)}{N(t)} + (1-\alpha) \cdot \frac{1 \times X_i(t)}{2E(t)} \quad ③$$

$$P(X_i(t+1) = X_i(t) + 1 \mid \vec{X}(t)) = 1 - ③$$

(Thm 7.1) for all $i \geq 1$,

$$\frac{X_i(t)}{t} \xrightarrow{\text{as } t \rightarrow \infty} C_i,$$

where $C_1 = \frac{\alpha}{\gamma + \alpha}$

$$\frac{C_i}{C_{i-1}} = \frac{\alpha + \frac{(1-\alpha)(i-1)}{2}}{1+\alpha + \left(\frac{1-\alpha}{2}\right)_i}, \quad i \geq 1$$

\Rightarrow degree $\eta \geq \frac{3}{2}$ implies
power law of t

(Note) (Initial States λ)

$$\frac{X_i(t)}{t} \xrightarrow{\text{as } t \rightarrow \infty} (\text{Init State}) C_i t^{-\beta}$$

weak form
 $(\alpha + \frac{3}{2}) \left(\frac{3}{2} \right)$

(Init) (Power Law of Init State)

$$C_j = C_1 \cdot \frac{C_2}{C_1} \cdot \frac{C_3}{C_2} \cdot \dots \times \frac{C_j}{C_{j-1}} = C_1 \prod_{i=2}^j \frac{C_i}{C_{i-1}}$$

$$\frac{C_j}{C_1} = \frac{\alpha + \frac{(1-\alpha)(j-1)}{2}}{1+\alpha + \left(\frac{1-\alpha}{2}\right)_j} = \frac{2\alpha + (1-\alpha)(j-1)}{2+2\alpha + (1-\alpha)_j} = 1 - \frac{2-j}{2+2\alpha + (1-\alpha)_j}$$

$$= 1 - \frac{1}{j} \frac{3-\alpha}{2+\alpha} + O\left(\frac{1}{j}\right)$$
$$= 1 - \frac{1}{j} \frac{3-\alpha}{2+\alpha} + O\left(\frac{1}{j^2}\right)$$

$$\log(C_j) = \log C_1 + \sum_{i=2}^j \log \left(1 - \frac{1}{j} \frac{3-\alpha}{2+\alpha} + O\left(\frac{1}{j^2}\right) \right)$$

$$\left(\log(1-x) \approx -x \quad \frac{x}{1+x} \approx 1-x \right)$$
$$e^x \approx 1+x \Leftrightarrow (e^x - 1) \approx x$$

$$\approx \log C_1 + \sum_{i=2}^j \frac{1}{j} \frac{3-\alpha}{2+\alpha}$$

$$\approx \log C_1 - \beta \log j, \quad \beta = \frac{3-\alpha}{2+\alpha}$$

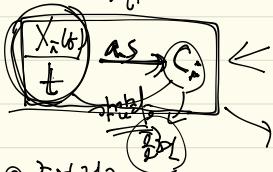
$$C_j \approx C_1 \cdot j^\beta \quad (j \geq \frac{3}{2})$$

initial state

* 1. 정의 2. 정리 3. 증명

① $P_A \rightarrow \text{node } \Rightarrow \text{stationary distribution}$ MC
degree distribution = 각 노드의 차수 / 노드 수에 대한 확률?
 $\frac{X_i(t)}{t} \rightarrow \text{stationary distribution}$
state transition dynamics

$X_i(t) \text{ st } X_{\bar{i}}(t) \text{ 대략 } \frac{1}{t} \text{ th } C_i \text{ 인가? } \text{ 증명?}$



② 증명

$$\frac{X_i(t)}{t} \rightarrow C_i$$

$$\frac{X_i(t)}{t} \leq \left| \frac{X_i(t)}{t} - C_i \right| + \left| C_i - \bar{X}_i(t) \right|$$

relative characterization of $\bar{X}_i(t)$

$$|X - E(X)| \leq M$$

A-H inequality

Thm 7.2

For all $\varepsilon > 0$, and for all $i \geq 1$

fix t_2 .

$$\frac{X_i(t)}{t} = \frac{C_i}{t} + o(t^\varepsilon)$$

$$\frac{X_i(t)}{t} = C_i + o(t^{\varepsilon-1}) \quad \xrightarrow[t \rightarrow \infty]{} 0$$

Lemma 7.3

$$P(|X_i(t) - \bar{X}_i(t)| \geq M) \leq 2 \exp\left(-\frac{M^2}{8t}\right) \quad (\text{A-H inequality})$$

Proof of Thm 7.1 A.S 3rd Borel-Cantelli Lemma

$$\left(\sum_{n=1}^{\infty} P_n \rightarrow 0 \right) \rightarrow X_m \xrightarrow{a.s.} X$$

이제 증명

$$\Pr \left(\left| \frac{X_i(t)}{t} - L_i \right| > \alpha(t) \right) \leq \Pr \left(\left| \frac{\bar{X}_i(t)}{t} - \frac{\bar{L}_i}{t} \right| + \left| \frac{\bar{X}_i(t)}{t} - L_i \right| > \alpha(t) \right)$$

\Rightarrow 여기 2번 째는 A(t)와 2번 째는 a(t)는 서로 같아요, \sum ∞ \Rightarrow $\Pr \left(\left| \frac{\bar{X}_i(t)}{t} - \frac{\bar{L}_i}{t} \right| > \alpha(t) \right)$ (Theorem 7.2)

$$= \Pr \left(\left| \frac{\bar{X}_i(t)}{t} - \frac{\bar{L}_i}{t} \right| > \alpha(t) - o(t^{-\varepsilon_1}) \right)$$

(Lemma 7.3)

$$\leq 2 \cdot \exp \left(- \frac{M^2}{8t} \right) = 2 \cdot \exp \left(- \frac{2\bar{L}_i^2 \cdot \alpha(t)^2}{8t} \right) = \left(\frac{2}{2t} \right)^2$$

$$\alpha(t) = \frac{4\sqrt{t} + 8t}{t} + o(t^{\varepsilon_1})$$

$$\begin{aligned} & t \rightarrow \infty \\ & \alpha(t) \rightarrow 0 \end{aligned}$$

$$\sum_{t=1}^{\infty} 2t^2 < \infty \Rightarrow \Pr \left(\left| \frac{\bar{X}_i(t)}{t} - L_i \right| > \alpha(t) \right) < \infty$$



$$\sum_t \frac{1}{t} = \infty, \sum_t \frac{1}{t^2} < \infty$$

증명

(Thm 7.2)

$$\Pr \left(\frac{\bar{X}_i(t)}{t} = X_i(t) \mid \text{모든 } (t-1) \right)$$

$$\Pr \left(\bar{X}_i(t) = X_i(t) \right)$$

(7.2)

$$\Pr \left(X_i(t+1) = X_i(t) \mid \bar{X}_i(t) \right)$$

$$= \Pr \left(\frac{X_i(t)}{N(t)} + (1-\lambda) \cdot \frac{1 \times X_i(t)}{2E(t)} \right)$$

$$\Pr \left(X_i(t+1) = X_i(t) + 1 \mid \bar{X}_i(t) \right) = 1 - \textcircled{1}$$

$$\mathbb{E}(X_i(t+1)) = \mathbb{E} \left(\mathbb{E}(X_i(t+1) \mid \bar{X}_i(t)) \right) = \sum_x \Pr(X_i(t)=x) \cdot \left(\underbrace{\left(\frac{(1-\lambda) \sum_{x=1}^{N(t)} x}{2E(t)} \right)}_{A(t)} + (x+1)(1-A(t)x) \right)$$

$$= \sum_x \Pr(X_i(t)=x) [x + 1 - A(t)x]$$

$$= \bar{X}_i(t) + 1 - A(t) \cdot \bar{X}_i(t)$$

where ~~여기~~

$$\boxed{\bar{X}_i(t+1) = \bar{X}_i(t) + 1 - A(t)\bar{X}_i(t)}$$

$\bar{X}_i(t)$ for a given t

$$\overline{X}_i(t+1) = ((1-\alpha_i t)) \overline{X}_i(t) + 1 \Rightarrow \overline{X}_i(t) = C_i t + o(t^\epsilon)$$

$$\Delta_i(t) = \overline{X}_i(t) - C_i t$$

$\boxed{\Delta_i(t+1)} \rightarrow \text{Homework} (\text{Homework 3})$

$$\Delta_i(t+1) = \Delta_i(t) \left(1 - \frac{\alpha}{N(t)} - \frac{1-\alpha}{2E(t)} \right) - C_i + 1 - C_i \left(\alpha + \frac{1-\alpha}{2} \right) + O(t^\epsilon)$$

$\alpha < 1 - \frac{1-\alpha}{2}$

$$\Delta_i(t+1) \leq \Delta_i(t) + O(t^\epsilon) \leq O(\log t) \Rightarrow \Delta_i(t) = O(t^\epsilon) \text{ for all } \epsilon > 0$$

$$C_i = \frac{2}{3t^\epsilon}$$

$$(i) \quad \overline{X}_i(t+1) = f(\overline{X}_i(t), \dots, \overline{X}_i(t)) \quad \boxed{\text{Homework}} \quad C_i, \quad \boxed{\frac{C_i}{C_{i+1}} \text{ small}}$$

Lemma 7.3 For all $i, t \geq 1$, and all $M > 0$,

$$P(|\overline{X}_i(t) - \underline{X}_i(t)| \geq M) \leq 2 \exp\left(-\frac{M^2}{8t}\right) \Rightarrow \text{Hoeffding inequality}$$

(i) Let i, t be fixed.

$$\overline{X}_i(t) = f(v(1), v(2), \dots, v(t))$$

where $v(s)$ is the node in $G(s-1)$ to which the node (i, U_S) attaches
new nodes

$$\text{Also, let } M(s) = E\left(\overline{X}_i(t) \mid v(1), \dots, v(s)\right), s=1, \dots, t \Rightarrow M(s-1) = E\left(\overline{X}_i(t) \mid v(1), \dots, v(s-1)\right)$$

$$M(0) = E[\overline{X}_i(t)] \Rightarrow \text{independent}(X)$$

(i) $M(s)$ is a martingale
(ii) $|M(s) - M(s-1)| \leq C$) \Rightarrow Hoeffding inequality

(i) easy? Doubt martingale

(ii) is true(?)

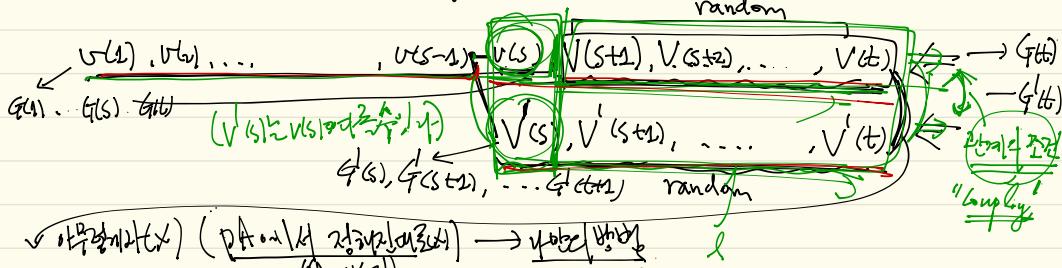
(ii) f is Lipschitz continuous $\frac{M(s)}{M(t)}$, $\Rightarrow \frac{M(s)}{M(t)} \leq L$ $M(s) - M(s-t) \leq L$ $Ls \leq t$

page 69: Corollary 6.4 (X_1, \dots, X_T independent) Coupling

Random variable

(X_1^D, X_2^D) deterministic
 (X_1^R, X_2^R) random

- Let s be fixed, and let the sequence $v(s), v(s+1), \dots, v(t)$ be given
- Let another random variable $V(s)$ of $G(s-1)$ be given, being distributed as the anchor node in $G(s)$, given $v(1), \dots, v(s-1)$



$$\checkmark \text{ (page 69) } (P(A \cap B)) = P(A)P(B) \rightarrow \text{Independent}$$

We will generate $\{V(s+1), \dots, V(t)\}$ $\{V(s), V(s+1), \dots, V(t)\}$, such that the following properties are satisfied:

(resp. $V(s+1), \dots, V(t)$)

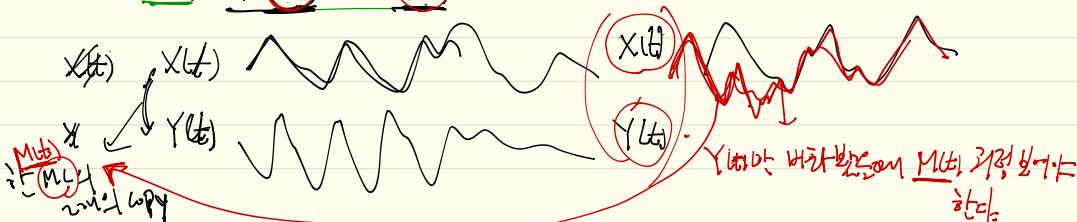
(page 82) (i) The distribution of $\{V(s+1), \dots, V(t)\}$ is that of the $(s+1)$ -th to t -th anchor nodes in the graph growth model that we consider, conditioned on the first s anchor nodes being $\{v(1), \dots, v(s)\}$

(resp. $v(1), v(2), \dots, v(s-1), V(s)$)

(ii) For all $l \in S, \dots, t$, and any nodes u in the node set $G(l), G(t)$,

the degree $d(u)$ of u in $G(l)$ coincides with $d(u)$ in $G(t)$,

unless $u = V(s)$ or $u = V(s)$



임의의 coupling은 확률분포?

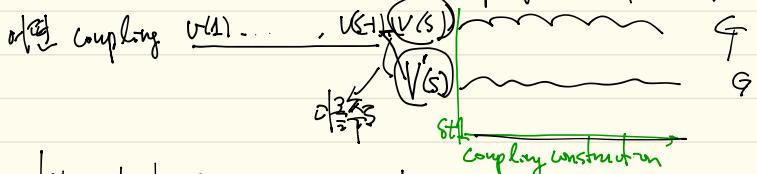
$$M(s) = f(v_{s1}, \dots, v_{sL}) = E\left[\left(\prod_{i=1}^L v_{si}\right) \mid v(1), \dots, v(s)\right]$$

$$\begin{aligned} & |M(s) - M(s-1)| \leq \epsilon \\ & = \sum P(V_{s+1}^t = v_{s+1}^t, V_s^t = v_s^t) \left[f(v_i^t) - f(v_i^{s+1}, v_s^t) \right] \\ & \quad \text{coupling} \quad \text{coupling} \quad \text{coupling} \\ & \quad (v(1), \dots, v(s-1)) \quad M(s-1) \quad v(s), v(s+1), \dots, v(t) \\ & \quad \text{coupling} \quad \text{coupling} \quad \text{coupling} \\ & \quad (v(s+1), \dots, v(t)) \quad V \leq 2 \\ & \quad \text{coupling} \quad \text{coupling} \quad \text{coupling} \\ & \quad (v(s+1), \dots, v(t)) \quad v(s), v(t) \end{aligned}$$

~~A~~ ≤ 2

Goal: (i) and (ii)를 만족하는 coupling을 찾는 것과 show

\Rightarrow how? Construction (Coupling 방법 | 조합법) (page 82의 half)



At each step $t=s+1, s+2, \dots, t$

Define a Bernoulli random variable $Y_t \in \{0, 1\}$ w.p. α and $1 - \alpha$ i.i.d.

attachment to a node uniformly at random
 if $(Y_1=0)$, choose an anchor node U (uniformly at random),
 $V(Q)=U \Rightarrow V'(Q)=U$

If $(Y_1=1)$, no horizontal attachment
 ① $U \notin \{V(S), V'(S)\}$

$$P(V(Q)=U) \mid Y_1=1, V_{S+1}^{l-1}, V_S^{l+1}) = \frac{d_{S+1}(U)}{2E(Q-1)} \text{ and}$$

$\boxed{V(Q) \Rightarrow V'(Q)}$

② $V(S) \text{ or } V'(S) \text{ are } \frac{\text{def}}{\text{def}}$ ($u, v \in \{V(S), V'(S)\}$)

$$P(V(Q)=U, V(Q)=v \mid Y_1=1, V_{S+1}^{l-1}, V_S^{l+1}) =$$

$\frac{d_{S+1}(U) \cdot d_{S+1}'(U)}{2E(Q-1) [d_{S+1}(V(S)) + d_{S+1}(V'(S))]}$

$$(i) P(V(Q)=U \mid V_{S+1}^{l-1}, V_S^{l+1}) = \frac{d}{N(Q-1)} + (1-d) \frac{d_{S+1}(U)}{2E(Q-1)}$$

$u \notin \{V(S), V'(S)\} \rightarrow \underline{\text{decreasing}} \text{ immediately true}$

$u \in \{V(S) \sim V \in V'(S)\}$

$$P(V(Q)=v \mid \underline{\quad}) = P(V(Q)=v \mid V(Q)=V(S)) + P(V(Q)=v \mid V(Q)=V'(S))$$

$$= \frac{d_{S+1}(V(S)) (d_{S+1}(V(S)) + d_{S+1}(V'(S)))}{2E(Q-1) (d_{S+1}(V(S)) + d_{S+1}(V'(S)))} = \frac{d_{S+1}(V(S))}{2E(Q-1)}$$

Similarly $\Pr(V(R) = V(s))$

Similarly $\Pr(V(q) = V(s))$, $\Pr(V(q) = V(s)) \dots$

(ii) $U \in \{V(s), V(s)\}$ degree concaves (凸凹)

