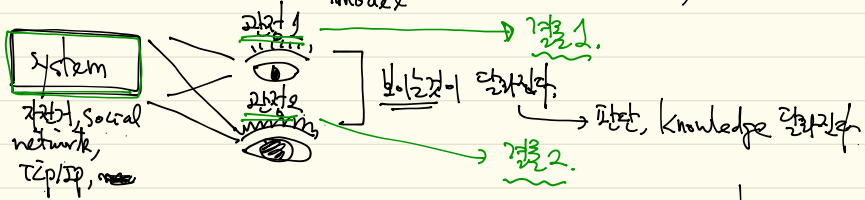


Lecture 16: (Chapter 5) From microscopic to macroscopic

↳ "fluid model", "mean-field approximation"

"fluid models for Internet"

(System (Dynamical system) model)



$$\{X_i\}, i=1, \dots, n$$

$$\frac{\sum_{i=1}^n X_i}{n}$$

관찰 1: $n \rightarrow \infty$ R.V. sum의 평균: $\frac{1}{n} \sum_{i=1}^n X_i$ stochastic 관측
 관찰 2: $n \rightarrow \infty$ R.V. sum의 분산: $\frac{1}{n^2} \sum_{i=1}^n X_i^2 - (\frac{1}{n} \sum_{i=1}^n X_i)^2$ deterministic 관측

→ 몬테카를로 (MC)

Microscopic

system size (변수/상태)에 따른 복잡성

1 step / 2 step modeling을 고려한다

Macroscopic

more abstract한 관점에서 "평균" over system의 여러 구성요소

approximation

관찰 1, 2의 관측을 이룬다
 (관찰 1, 2)

(Q) $\frac{1}{n} \sum_{i=1}^n X_i$. Microscopic model로부터

Macroscopic model로 접근(?)



→ $\bar{x}(t) = f(x(t))$ 1. what? 2. 수학적 모델
 "평균" → 수학적 모델링을 통한 접근

Example

- ① Poisson process
- ② M/M/1 Queue
- ③ SIS epidemics

(Macro)
 (Macro)

③ $P_r (X_m(t+h) = \bar{x}+1 \mid X_m(t) = \bar{x}) = h \cdot \frac{\beta}{m} \bar{x}(m-\bar{x}) + o(h)$ microscope

$P_r (X_m(t+h) = \bar{x}-1 \mid X_m(t) = \bar{x}) = \delta \bar{x} h + o(h)$

$P_r (X_m(t+h) = \bar{x} \mid X_m(t) = \bar{x}) = 1 - h \left(\frac{\beta}{m} \bar{x}(m-\bar{x}) + \delta \bar{x} \right) + o(h)$ (d) $\frac{d}{dt}$

$X_m(t)$: total number of individuals

$x(t)$: relative number of individuals

$x(t) = \beta x(1-x) - \delta x$

Macroscopic

fluid model mean-field approximation approximation

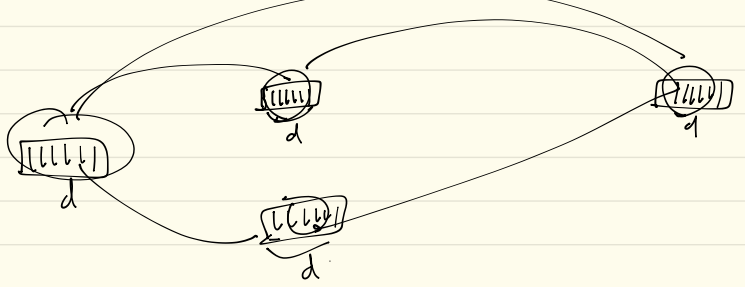
show

Kurtz Theorem (Chapter 5.3):

Markov jump process (Poisson process $\lambda \frac{1}{2}$ $\frac{1}{2}$ $\frac{1}{2}$, Birth/Death $\frac{1}{2}$ $\frac{1}{2}$)

M/M/1 queue

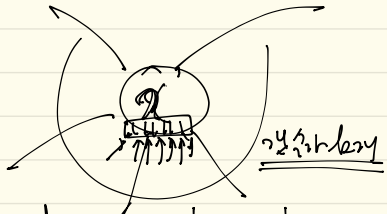
- Population of individuals belonging to "d" different species
-
- d=3
- $X_i(t)$, $i=1, \dots, d$: total # of individuals of species i at time t
 - $X(t) = (X_i(t) = i=1, \dots, d)$:



- We say that $(X(t))$ is Markov jump process when (x) (y)
 for any given two states $x, y \in \mathcal{N}^d$, the process jumps from x to y
 upon expiration of a random timer with exponential distribution
 whose parameter depends on x and y



- Assume that \exists a finite # of possible jump directions
 $(e_i)_{i=1, \dots, k}$



$(\frac{2D}{2})$ best def \mathcal{N}^d ? \mathcal{N}^d

possible # of jump directions $\leq k$

(cf) when e_i 's entries are restricted to $\{d, 0, 1\}$

(cm) M/M/1

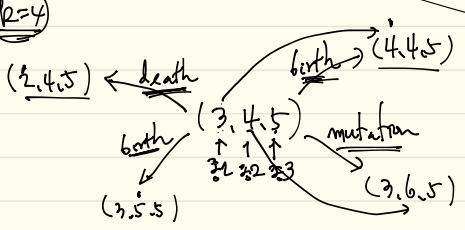
\Rightarrow Birth-and-death process

(cf) Poisson process $e_i \rightarrow \dots$

$e_i \in \mathcal{N}^d$

Let us denote by $\lambda_i(x)$ "the rate from (x) to $(x+e_i)$ "

(EX) $(d=3, k=4)$



\rightarrow exponential time ξ_i
 \rightarrow $\frac{1}{\lambda_i}$

$(\frac{h}{2})$: scaling parameter (m) ($\frac{h}{2} \rightarrow 0, m \rightarrow \infty$)

For any fixed (m) \longrightarrow MJP (m) , $X_m(t)$

$n=1 \longrightarrow X_1(t)$

$2 \longrightarrow X_2(t)$

\vdots

$n \longrightarrow X_n(t)$

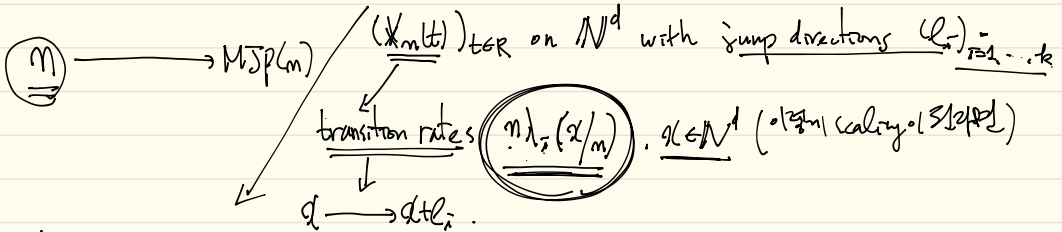
\vdots

$\infty \longrightarrow X_\infty(t)$

Question $(X_\infty(t))$ is (어떻게) 이동할 것인가?
(어떻게) 정당 \longleftarrow 정

Theorem 5.2 \longrightarrow Theorem 5.3

Theorem 5.3 (Kurtz's Theorem) (page 67 = 어떻게)



Then, $\{X_m(t)\}_{t \geq 0}$ satisfies:

$$\left(\begin{aligned} \Pr(X_m(t+h) = q + e_i \mid X_m(t) = q) &= \eta_i \cdot h \cdot \lambda_i(x/m) + o(h), \quad q \in \mathbb{N}^d, i=1, \dots, k \\ \Pr(X_m(t+h) = q \mid X_m(t) = q) &= 1 - m \cdot h \sum_{i=1}^k \lambda_i(x/m) + o(h), \quad q \in \mathbb{N}^d \end{aligned} \right)$$

• Let the function $\bar{\pi}: \mathbb{R}^d \rightarrow \mathbb{R}^d$ be defined by e_1, \dots, e_k

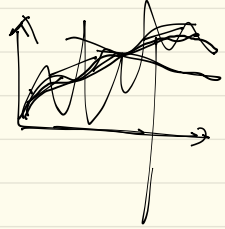
$$\Rightarrow \bar{\pi}(x) := \sum_{i=1}^k \frac{e_i \cdot \lambda_i(x)}{\eta_i} = \underbrace{(\uparrow, \uparrow, \uparrow, \dots, \uparrow)}_{d \text{ dim}}$$

• Let $\bar{e} := \max_{i=1, \dots, k} \{e_i\}$. \nearrow some arbitrary (norm) distribution

Assume the following conditions hold,

→ C1: $\bar{\lambda} := \max_{i=1, \dots, k} \sup_{x \in \mathbb{R}^d} \lambda_i(x)$ is finite. ← (어떻게 보느냐?)

→ C2: The function \bar{F} is Lipshitz-continuous, i.e., \exists a constant M s.t.
 $|\bar{F}(x) - \bar{F}(y)| \leq M|x-y|$, for all $x, y \in \mathbb{R}^d$
 $\frac{|\bar{F}(x) - \bar{F}(y)|}{|x-y|} \leq M$



→ C3: $\lim_{m \rightarrow \infty} \frac{1}{m} X_m(t) = \alpha(t)$ a.s. (어떻게 보느냐?)

Let $q: \mathbb{R}_+ \rightarrow \mathbb{R}^d$ be the solution of the following equation: $\overset{\text{integral}}{\int} \dot{q}(t)$

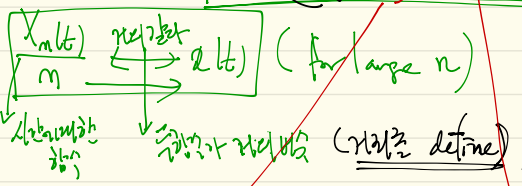
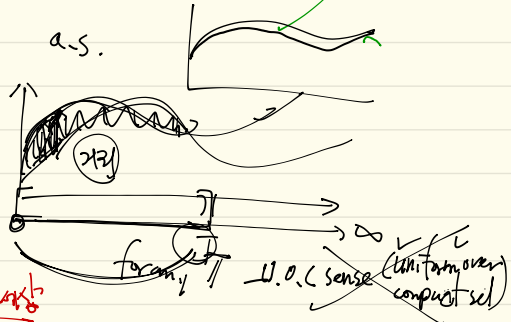
$$\overset{\text{deterministic}}{\text{Macroscopic}} \left[\frac{X_m(t)}{m} \right] \rightarrow q(t) = q(0) + \int_0^t \bar{F}(q(s)) ds$$

then, we have: for any fixed $\varepsilon, \delta > 0$, for sufficiently large n (deterministic)

$$P_r \left(\sup_{0 \leq t \leq T} \left| \frac{X_m(t)}{m} - q(t) \right| \geq \varepsilon \right) \leq 2k \exp \left(-m T \bar{\lambda} h \left(\frac{\varepsilon e^{-hT}}{2kT \bar{\lambda} \varepsilon} \right) \right)$$

where $h(t) = (1+t) \log(1+t) - t$. Moreover,

$$\lim_{m \rightarrow \infty} \left(\sup_{0 \leq t \leq T} \left| \frac{X_m(t)}{m} - q(t) \right| \right) = 0 \text{ a.s.}$$



Kurtz theorem: macroscopic → $\text{deterministic functional version}$

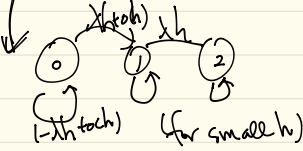
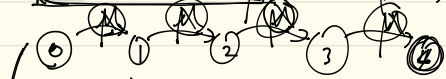
"Functional Law of Large Numbers"
 Central Limit Theorem

Lecture 11 (part 2) : Kurtz Theorem $\frac{2}{2}$

- ① Poisson process
- ② SIS epidemics

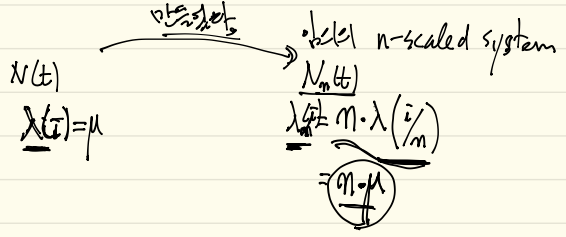
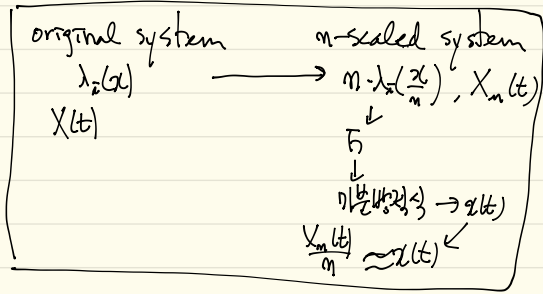
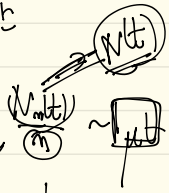
< Poisson process >

$N(t)$ with rate λ (Birth process)



$N(t) \approx \square$
deterministic system, λt

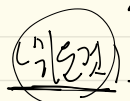
By Kurtz theorem,



$\dot{x}(t) = f(x(t)), f(x) = \mu$
 $= \mu \rightarrow x(t) = \mu t$

$\lim_{n \rightarrow \infty} \left(\sup_{0 \leq t \leq T} |N(t) - \mu t| \right) = 0 \text{ a.s. } \left(N(t) \approx \mu t \right)$

Proposition 5-2



$\frac{N(t)}{n} \approx \frac{N_n(t)}{n} = \frac{N(t)}{n}$

$N(t) \rightarrow$ time on $\frac{1}{n}$ scale
space on $\frac{1}{n}$ scale

$\frac{N(t)}{n} \approx x(t)$

$\frac{N(t)}{n} \approx \frac{N(t)}{n} \rightarrow \mu \cdot \frac{N(t)}{n}$

"Functional law of large n 's"

$\frac{N(t)}{n} \rightarrow \mu \cdot \mu t$

$E(N(t)) = \mu t$

< SIS epidemics

$$\Pr(X_m(t+h) = \bar{i}+1 \mid X_m(t) = \bar{i}) = h \left(\frac{\beta}{m} \bar{i}(m-\bar{i}) + o(h) \right)$$

$$\Pr(X_m(t+h) = \bar{i}-1 \mid X_m(t) = \bar{i}) = \delta \bar{i} h + o(h)$$

$$\Pr(X_m(t+h) = \bar{i} \mid X_m(t) = \bar{i}) = 1 - h \left(\frac{\beta}{m} \bar{i}(m-\bar{i}) + \delta \bar{i} \right) + o(h)$$

dim setup of "n" nodes scaling parameter β
 MCMC system



$$\bar{i} \rightarrow \bar{i}-1 : \delta \bar{i} h = m \cdot \lambda^-(\bar{i}/m)$$

$$\bar{i} \rightarrow \bar{i}+1 : \frac{\beta}{m} \bar{i}(m-\bar{i}) = m \cdot \lambda^+(\bar{i}/m)$$

what is $\lambda^+(\bar{i})$? $m \cdot \lambda^+(\frac{\bar{i}}{m}) = \delta \bar{i} h = m \cdot \delta (\frac{\bar{i}}{m}) \cdot h \Rightarrow \lambda^+(\bar{i}) = \delta \bar{i}$

$\lambda^-(\bar{i})$? $m \cdot \lambda^-(\frac{\bar{i}}{m}) = \frac{\beta}{m} \bar{i}(m-\bar{i}) = m \cdot \beta (\frac{\bar{i}}{m}) (1 - \frac{\bar{i}}{m}) \Rightarrow \lambda^-(\bar{i}) = \beta \bar{i}(1-\bar{i})$

$\bar{h} \rightarrow \alpha$

$e_+ = 1, e_- = -1$

$$\bar{h}(\bar{i}) = e_+ \lambda^+(\bar{i}) + e_- \lambda^-(\bar{i}) = \beta \bar{i}(1-\bar{i}) - \delta \bar{i}$$

$$\bar{h}(\alpha) = \beta \alpha(1-\alpha) - \delta \alpha$$

first simplification, $\delta = 1, \alpha(0) = y$

$$\bar{q}(t) = \bar{h}(\alpha(t)) = \beta \alpha(1-\alpha) - \delta \alpha$$

homework

$$\alpha(t) = \frac{(1-\beta)y e^{(\beta-1)t}}{(1-\beta) - \beta y (1 - e^{(\beta-1)t})}$$

$\frac{X_m(t)}{m} \sim \alpha(t)$

no \rightarrow $\frac{X_m(t)}{m}$ \rightarrow $\alpha(t)$ model "portion" "fraction" $\alpha(t) = \frac{X_m(t)}{m}$

If $\beta > 1, \alpha(t) \rightarrow (1 - \frac{1}{\beta})$ as $t \rightarrow \infty$

If $\beta < 1, \alpha(t) \rightarrow 0$ as $t \rightarrow \infty$

Lecture 11 (part 3) Kurtz Theorem of 2006

Proposition (5.2) (Kurtz Theorem of ^(SE version) poisson process version)

Let $N(t)$ be a poisson process with unit rate. Then for any $\epsilon > 0$ and $T > 0$,

$$\Pr\left(\sup_{0 \leq t \leq T} |N(t) - t| \geq \epsilon\right) \leq e^{-\lambda \epsilon^2 / T}, \quad h(t) = (t+t) \log(t+t) - t.$$

(Pf) Doob's inequality: $N(t) - t$, $t - N(t)$ are submartingales

$$\Pr\left(\sup_{0 \leq t \leq T} |N(t) - t| \geq \epsilon\right) \leq \Pr\left(\sup_{0 \leq t \leq T} N(t) - t \geq \epsilon\right) + \Pr\left(\sup_{0 \leq t \leq T} t - N(t) \geq \epsilon\right)$$



① \rightarrow ① \hookrightarrow for any $\theta > 0$ \rightarrow ②

$$\Pr\left(\sup_{0 \leq t \leq T} (N(t) - t) \geq \epsilon\right) \leq \Pr\left(\sup_{0 \leq t \leq T} e^{\theta(N(t) - t)} \geq e^{\theta \epsilon}\right) \leq \frac{E\left(e^{\theta(N(T) - T)}\right)}{e^{\theta \epsilon}}$$

\Rightarrow Why $N(t) - t$ submartingale?

(Continuous version of Martingale definition)
 $\{X(t)\}_{t \geq 0}$ is a submartingale if and only if the following holds:

$$E\left(X(t) \mid \{X(u)\}_{u \leq s}\right) \geq X(s), \quad 0 \leq s \leq t$$

$$\begin{aligned} E\left(N(t) - t \mid \{N(u) - u\}_{u \leq s}\right) &= E\left(\underbrace{N(t-s) + N(s)}_{\substack{\text{independent increments} \\ \text{submartingale } N(s) - s}} - (t-s) - s \mid \{N(u) - u\}_{u \leq s}\right) \\ &= E\left(N(t-s) - (t-s) \mid \emptyset\right) + E\left(N(s) - s \mid \emptyset\right) \\ &= E\left(N(t-s) - (t-s)\right) + E\left(N(s) - s\right) \\ &= 0 + 0 = 0 \end{aligned}$$

$$\frac{E(e^{(\epsilon T - T)\theta})}{e^{\theta \epsilon}} = e^{\theta \epsilon} \cdot e^{-T\theta} E(e^{N(T)\theta})$$

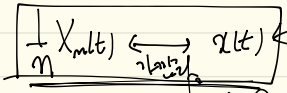
$$= e^{\theta \epsilon} \cdot e^{-T\theta} \cdot e^{T(e^{\theta} - 1)} = e^{(-\theta(\epsilon T) + T(e^{\theta} - 1))}$$

$N(T) \sim$ Poisson random variable
with parameter T

$$\beta^* = \log(1 + \epsilon/T)$$

when $\theta \rightarrow 0^+$
 $\approx e^{-T\beta(\epsilon/T)}$, $h(t) = \dots$
substituting β^*

(Proof)

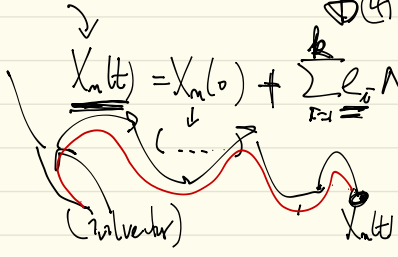


initial value solution

$X_m(t)$: Markov system

- $X_m(t)$ dynamics is described by $X_m(t)$ of some character.
- ~~$X_m(t) =$~~ (?) (jump direction)
- with $X_m(t)$ is a 1 ISB, h is a 1 ISB.

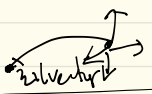
d-dimensional vector



- obtain $\alpha(t) \approx$

$$X_m(t) = X_m(0) + \sum_{i=1}^k \lambda_i N_i \left(\int_0^t \underbrace{\mu_i(s)}_{\text{"rate"}} ds \right)$$

N_i : independent unit-rate poisson process
i.e. N_1, N_2, \dots, N_k (unit rate poisson clock λ_i kth)



$$\frac{X_m(t)}{m} = \frac{X_m(0)}{m} + \sum_{i=1}^k \lambda_i \frac{N_i}{m}$$

Define $\bar{N}_i(t) = N_i(t) - \lambda_i t$ "centered Poisson process"

Centrality \rightarrow $E(\bar{N}_i(t)) = 0$

$$\frac{X_m(t)}{m} = \frac{X_m(t_0)}{m} + \frac{1}{m} \left(\sum_{i=1}^k \frac{L_i}{\Gamma_i} \left(\overline{X}_i \left(\int_0^t m \cdot X_i \left(\frac{X_m(s)}{m} \right) ds \right) \right) \right) + \frac{1}{m} \sum_{i=1}^k \int_0^t m \cdot X_i \left(\frac{X_m(s)}{m} \right) ds$$

Let $X_m(t) = \frac{X_m(t)}{m}$. Recall that $\Gamma_i(x) = \sum_{j=1}^k L_j \cdot x_j(x)$

$$X_m(t) = X_m(t_0) + \sum_{i=1}^k \frac{L_i}{m} \left(\overline{X}_i \left(\int_0^t \underbrace{\sum_{j=1}^k L_j X_j \left(\frac{X_m(s)}{m} \right) ds}_{\Gamma_i \left(\frac{X_m(s)}{m} \right)} \right) \right) + \int_0^t \underbrace{\sum_{i=1}^k L_i X_i \left(\frac{X_m(s)}{m} \right) ds}_{\Gamma_i \left(\frac{X_m(s)}{m} \right)}$$

기함/기함: $|X_m(t) - x(t)|$

정확도/정확도: $x(t) = x(t_0) + \left(\int_0^t \Gamma(x(s)) ds \right)$

$$\begin{aligned} |X_m(t) - x(t)| &\leq |X_m(t_0) - x(t_0)| + \left| \int_0^t \left(\Gamma(X_m(s)) - \Gamma(x(s)) \right) ds \right| \\ &\leq M \left(\int_0^t |X_m(s) - x(s)| ds \right) + \left| \int_0^t m \cdot X_i \left(\frac{X_m(s)}{m} \right) ds \right| \\ &\quad + \sum_{i=1}^k \left(\frac{L_i}{m} \left| \overline{X}_i \left(\int_0^t \left(\frac{X_m(s)}{m} - x(s) \right) ds \right) \right| \right) \end{aligned}$$

$$|f(x)| \leq a + \int_0^t |f(s)| ds + \dots$$

$$|f(x)| \leq \Delta + \int_0^t |f(s)| ds + \Delta \Rightarrow f(x) \leq \Delta$$

$f(x) = \int_0^t f(x) = \int_0^t f(x) ds$

$f(x) = \int_0^t \int_0^t f(s) ds$

$f(x) = \Delta$

Proposition 5.6 (Gronwall's Lemma)

Let u be a bounded real-valued function on $[0, T]$. s.t.

$u(t) \leq a + \int_0^t u(s) ds$, for all $t \in [0, T]$

then $u(t) \leq a e^{bt}$ and $u(T) \leq a e^b$

$$\sup_{0 \leq t \leq T} |Y_m(t) - X(t) - M \int_0^t |Y_m(s) - X(s)| ds| \leq |Y_m(0) - X(0)| + \sup_{0 \leq t \leq T} \left(\sum_{i=1}^m |K_i| \frac{1}{N_i} \int_0^t m_i |Y_m(s)| ds \right)$$

for sufficiently large N (normalised)

$$\Pr \left(\sup_{0 \leq t \leq T} |Y_m(t) - X(t)| \geq 2\varepsilon \right) \leq \Pr \left(\sup_{0 \leq t \leq T} \left(\sum_{i=1}^m |K_i| \frac{1}{N_i} \int_0^t m_i |Y_m(s)| ds \right) \geq \varepsilon \right)$$

Proposition 2.2 general version
 $\leq 2K e^{-mT \bar{\lambda} h \left(\frac{\varepsilon}{KTE\lambda} \right)}$

$$\Pr \left(\sup_{0 \leq t \leq T} |Y_m(t) - X(t)| \geq 2\varepsilon \right) \leq 2K e^{-mT \bar{\lambda} h \left(\frac{\varepsilon}{KTE\lambda} \right)}$$

$$\leq \Pr \left(\sup_{0 \leq t \leq T} \left(|Y_m(t) - X(t)| - M \int_0^t |Y_m(s) - X(s)| ds \right) \geq 2\varepsilon \right) \stackrel{b=M}{a=2\varepsilon} \leq 2K e^{-mT \bar{\lambda} h \left(\frac{\varepsilon}{KTE\lambda} \right)}$$

$$\Pr \left(\sup_{0 \leq t \leq T} |Y_m(t) - X(t)| \geq \varepsilon \right) \leq 2K e^{-mT \bar{\lambda} h \left(\frac{\varepsilon}{KTE\lambda} \right)}$$

$$\Rightarrow \lim_{m \rightarrow \infty} \sup_{0 \leq t \leq T} |Y_m(t) - X(t)| = 0 \text{ a.s.} \quad \int f_m < \infty$$

Borel-Cantelli Lemma Let $A_m(\varepsilon) = \{ |X_m - X| \geq \varepsilon \}$
 Consider a seq. of r.v.s $\{X_m\}$
 If $\sum_m P(A_m(\varepsilon)) < \infty$ for all ε , then $X_m \rightarrow X$ a.s.

a.s. $\xrightarrow{0} \xrightarrow{(\otimes)}$ convergence in probability (Q) under what condition ???
 c.i.p. \rightarrow a.s.

$\forall \epsilon > 0 \quad \Pr(A_n(\epsilon)) \xrightarrow{n \rightarrow \infty} 0$ (convergence in probability)

• $\forall \epsilon > 0 \quad \sum_{n=1}^{\infty} \Pr(A_n(\epsilon)) < \infty \rightarrow \underline{a.s.} \quad \text{wz} \quad \sum a_n < \infty$
 $\downarrow \quad \uparrow$
 $\Pr(A_n(\epsilon)) \xrightarrow{n \rightarrow \infty} 0$ $\lim_{n \rightarrow \infty} a_n = 0$

B-C Lemma \Rightarrow $X_n \rightarrow X$ a.s. sufficient (sufficient)

① $\sum \Pr(A_n(\epsilon)) < \infty$

② $\sum_{n=1}^{\infty} \Pr(A_n(\epsilon)) < \infty \rightarrow \Pr(A_n(\epsilon)) \xrightarrow{n \rightarrow \infty} 0$

$\Pr(|X_n - X| > \epsilon) \Rightarrow \sum_{n=1}^{\infty} \frac{1}{n^2} < \infty$ (a.s.)
 $\Rightarrow \sum \frac{1}{n^2} < \infty$
 $= \frac{1}{9^2} < \infty \quad \sum \frac{1}{n^2} < \infty$

$\frac{x}{|z_n|}$

Gronwall's Lemma \Rightarrow Converge in probab. (a.s.) \Rightarrow a.s. (X)
Mode. f. Convergence

$\frac{1}{n^2}$

B.C. test $\left[\sum \Pr(A_n(\epsilon)) < \infty \right] \Rightarrow \text{a.s.}$